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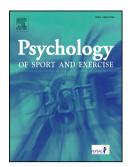
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Are trait self-control and self-control resources mediators of relations between executive functions and health behaviors?

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Contributors

C.F., M.d.C and T.D. designed the study protocol. C.F. designed the analyses. M.D. C.F, M.d.C, R.B, and T.D. collected the data. C.F. analyzed the data. C.F., M.d.C and T.D. drafted the manuscript. A.C., R.B, M.B, and B.C. critically appraised the manuscript and significantly contributed to its content and structure. All authors approved the final version of the manuscript.

Data sharing

Data and code for the analysis are available on OSF at

https://osf.io/hpsjw/?view_only=6e28c8307294494e9eec45d2670efd8d

1

Abstract

2 This study investigated associations between executive functions (i.e., inhibition, working 3 memory, cognitive flexibility) and individual differences in self-control and health behaviors. We examined whether executive functions predict physical activity, sedentary activity, and 4 5 healthy and unhealthy diets, and whether trait self-control and self-control resources mediate 6 these associations. Three hundred and eighty-five participants completed a questionnaire 7 assessing trait self-control and self-control resources, physical activity, sedentary activity, and 8 healthy and unhealthy diets. They also performed three randomly ordered cognitive tasks, a 9 stop-signal task (i.e., inhibition), a letter memory task (i.e., updating), and a number-letter 10 task (i.e., switching). Structural equation modeling revealed that self-control resources predicted positively physical activity ($R^2 = .08$), negatively sedentary activity ($R^2 = .03$) and 11 positively healthy diet ($R^2=10$). Moreover, trait self-control predicted positively healthy diet 12 $(R^2 = .10)$ and negatively unhealthy diet $(R^2 = .19)$. Moreover, analyses revealed that 13 switching significantly predicted self-control resources, and highlighted three totally mediated 14 15 relations between this executive function and physical activity, sedentary activity and healthy 16 diet. However, no evidence was found supporting associations between inhibition and updating, and health behaviors, or relations mediated by self-control for these executive 17 functions. The findings suggest the importance of trait self-control and self-control resources 18 19 for health behavior adoption and pave the way for studies exploring the role of the executive 20 functions in an affective context. Open materials [https://osf.io/hpsjw/].

21

22 Keywords: self-control resources, trait self-control, inhibition, updating, switching, health 23 behaviors

1

1. Introduction

2 According to the World Health Organization, physical inactivity and an unhealthy diet are 3 among the most important risk factors for noncommunicable diseases, causing one death 4 every seven seconds and one death every three seconds, respectively (Forouzanfar, Afshin, 5 Alexander, Biryukov, et al., 2016; Lee et al., 2012). Changing these unhealthy behaviors by 6 improving the regularity of physical activity and healthiness of diet could prevent 16 million 7 premature deaths each year (Forouzanfar, Afshin, Alexander, Anderson, et al., 2016). Despite 8 widespread declarations of intention to adopt healthy behaviors, most people fail to reach 9 minimum recommendations (Ford et al., 2011). In this context, some promising conflict resolution models could effectively promote health behaviors (Sniehotta et al., 2014). 10 11 Conflict resolution models (e.g., the integrative self-control theory, Kotabe et 12 Hofmann, 2015; the goal-conflict model, Stroebe, 2022) emphasize that health-behavior 13 facilitators (e.g., health-behavior goals, such as an intention to do more physical activity) face 14 barriers (e.g., temptation toward the competing behavior, such as a desire to engage in a 15 sedentary activity), leading to motivational conflicts that need to be resolved (Forestier, et al., 2022a; Rabiau et al., 2006). An adaptative resolution of this conflict (i.e., one in favor of the 16 17 goal) promotes the health-behavior goal (e.g., going for a run, snacking on an apple). A 18 maladaptive one promotes the competing behavior (e.g., remaining on the couch, snacking on 19 a chocolate bar) (e.g., Gillebaart et al., 2016). Here, self-control is defined as the self-20 regulation operationalization, by which an individual resolves a motivational conflict, through 21 either an effortless or effortful strategy (Forestier et al., 2022b; Gillebaart, 2018). As such, it 22 is expected to play a key role in promoting the resolution of motivational conflicts, favoring 23 healthy behaviors. Indeed, some studies showed correlations between self-control and healthy 24 diet and weight control with a small-to-medium effect size (de Ridder et al., 2012, r = .17), 25 and self-control and physical activity (Pfeffer & Strobach, 2018, r = .29).

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In 2018, Forestier et al. identified individual differences in trait self-control and self-26 27 control resources, as two dimensions independently associated with health behaviors. Despite 28 their importance, self-control resources have been mostly considered in ego-depletion 29 research (e.g., Rouse et al., 2013; see Forestier et al., 2022 for more details) and rarely on research on self-control and health behaviors (de Ridder et al., 2018). However, defined as 30 "the objective and subjective amounts of energy available for the self to initiate a self-control 31 32 act" (Forestier et al., 2022b, p. 21), the self-control resources seem crucial for motivational 33 conflict resolution, and can be assessed through subjective perception of energy availability, 34 or physiological markers (e.g., vagal activity). Accordingly, trait self-control distinguishes 35 individuals with a more or less tendency to successfully resolve motivational conflicts. In 36 parallel, individual differences in self-control resources distinguish individuals with a more or 37 less tendency to experience a high level of self-control resources (Forestier et al., 2018). 38 Authors identified that a remarkable portion of variance (40%) in self-control resources is 39 found at the between-person level (Smolders et al., 2013). In short, individuals with high trait 40 self-control consume a healthier diet than those with low trait self-control; individuals with 41 high self-control resources practice more physical activity, are less sedentary, and eat 42 healthier than individuals with low self-control resources. These results are in line with Finne 43 et al. (Finne et al., 2019), highlighting that within-person self-control fluctuations are more 44 important than between-person differences to predict physical activity. Moreover, a recent 45 study showed that within-person variations of self-control strength (i.e., a scale that could 46 refer to what Forestier et al., (2022) proposed as "self-control willingness") predict physical 47 activity but not trait self-control (Englert & Rummel, 2016).

48 Despite these interesting results, Forestier et al. (2018) did not specifically examined 49 the predictors of individual differences in the two self-control dimensions. Some hypotheses 50 on the executive functions have been already put forward that can partly explain the

51 differences in these dimensions (e.g., Hofmann et al., 2012, Table 2). Specifically, inhibition, 52 updating and switching (also called "inhibitory control", "working memory", and "cognitive 53 flexibility" respectively) is likely related to self-control, and individuals with higher executive 54 abilities would, therefore, be good self-controllers (Friese et al., 2011; Hofmann et al., 2011). Moreover, executive functions could operationalize the "self-control capacity" that makes 55 56 self-control success possible (Forestier, et al., 2022b; Kotabe & Hofmann, 2015). Presumably, 57 for instance, individuals with high inhibition capacity should present high self-control through 58 a better ability to override prepotent responses such as habits or impulses that are 59 incompatible with goals attainment. Likewise, individuals with high working memory 60 abilities could better maintain information in an active, quickly retrievable state and shield this information from distraction. They are therefore better able to use goal-relevant 61 62 information. Finally, relative to individuals with lower switching ability, individuals with high 63 switching may present high self-control through a high flexibility (opposed to "rigidity"), 64 which allow goal attainment by abandoning suboptimal means and selecting alternative 65 means to reach the same goal (see Hofmann et al., 2011, 2012, for full theoretical discussion). Additionally, structurally and functionally, brain networks in the prefrontal cortex that control 66 67 executive function, overlap with networks involved in self-control (e.g., right inferior frontal 68 gyrus, Cipolotti et al., 2016; Lopez et al., 2016).

Several assumptions on how executive functions and self-control might interact have
been proposed, including the mediation model hypothesis (Hofmann et al., 2012). To date,
most empirical self-control studies examined executive functions as moderators of the relation
between self-control and health behaviors (Hofmann et al., 2009a; Pfeffer & Strobach, 2017),
with mitigated results. Hence, we currently aimed at addressing the rarely explored approach
of executive functions as direct predictor of self-control. As far as we know, only Saunders et
al. (2018) and Necka et al. (2018) have investigated the association between executive

functions and self-control. Saunders et al. (2018) found no evidence of a correlation between 76 77 trait self-control and inhibition. Using structural equation modeling, Necka et al. (2018) found no significant association between a "trait self-control" latent variable and an "executive 78 79 functions" latent variable. However, methodological improvement would allow some results-80 related pitfalls to be avoided. First, only the relations between latent variables were tested, 81 without examining the independent contribution of each executive function to trait self-82 control. Second, these previous studies never focused on self-control resources, which are 83 correlated with trait self-control but remain an independent dimension to be considered 84 (Forestier et al., 2018). Third, the possibility that trait self-control and executive functions are 85 associated with health behaviors was not investigated. Yet, the relation between executive 86 functions and health behaviors deserves attention. Indeed, another study showed that 87 individuals with the highest inhibition adopted a less unhealthy diet (Hofmann et al., 2009). 88 Similarly, high updating abilities have been associated with more physical activity 89 (Lambourne, 2006; Pfeffer & Strobach, 2017). Finally, a bi-directional relationship between 90 executive functions and health behavior has been proposed, with individuals with high 91 executive abilities being more likely to adopt a healthy lifestyle that would, in turn, enhance 92 their executive functions in the long run (Allan et al., 2016). Empirical data has recently been 93 reported supporting this relation (Cheval et al., 2020). Accordingly, if executive functions are 94 correlated with self-control, as advanced theoretically (Hofmann et al., 2012), they could 95 promote health behaviors through direct and indirect effects, partially mediated by trait self-96 control and self-control resources. Such assumptions have never vet been tested.

97 The current study aimed to investigate the independent contribution of each executive 98 function to self-control and four crucial health behaviors to prevent noncommunicable 99 diseases, while considering differences among individuals in their trait self-control and self-100 control resources. For these purposes, structural equation models were used, one for physical

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101 activity versus sedentary activity, and a second for a healthy versus an unhealthy diet, as in 102 Forestier et al., (2018). In line with theoretical discussions (Hofmann et al., 2012), we 103 hypothesized that executive functions would positively predict trait self-control and self-104 control resources (H1) and healthy behaviors (physical activity and healthy diet) (H2), and 105 negatively predict unhealthy behaviors (sedentary activity and unhealthy diet) (H3). Similarly, 106 we hypothesized that trait and self-control resources would positively predict healthy 107 behaviors (H4) and negatively predict unhealthy behaviors (H5). Finally, the mediated 108 relation implies that (H6) executive functions will be positively related to trait self-control 109 and self-control resources, which will in turn be related to health behaviors¹. Figure 1 110 summarizes our hypotheses.

111

2. Method

112 Overview

113 Participants were recruited via social media, personal mailing lists, and direct advertising messages during classes. They were all students at sports and psychology faculties 114 115 of three different universities. Data were collected over three weeks (November 2021) by 116 completion of three cognitive tasks and four questionnaires on Inquisit web version 6.3.2.0 117 (Computer software) (data hosted by Inquisit, Europe repository) (during a single session 118 lasting 1h15). All procedures in this study complied with APA ethical principles. Informed 119 consent was obtained from all participants before the beginning of the study. It should also be 120 noted that the participants were informed that the online study was anonymous and 121 confidential: only a self-generated code allowed their identification.

¹No differential hypotheses regarding the role of trait self-control and self-control resources were formulated because of the rare existing evidences about relations between self-control resources and health behaviors. Moreover, we did not formulate differential hypotheses regarding the role of self-control and behaviors, because (a) despite differences, health behaviors considered share similarities (e.g., long-term benefits), as well as unhealthy behaviors share some (e.g., appetitive) (McEachan et al., 2010); and (b) literature stressed that self-control behave in comparable manner on different healthy (e.g., positive relation) and unhealthy behaviors (e.g., negative relation) (de Ridder et al., 2012).

122 Participants and sample size

123 Three procedures were used for estimating our sample size, and its relevance. The first 124 procedure aims to estimate a minimal sample size to reach. The second one aims to determine 125 a stopping rule to fix when data collection will end. The last one aims to compare the sample 126 size reached and the smallest detectable effect size with a certain power to what the literature 127 identified in past studies with common features, to consider whether another phase of data 128 collection is necessary or not. Precisely, first, we used a method specific to structural equation 129 modeling (MacCallum et al., 1996, 2006a; MacCallum & Hong, 1997) to estimate an a priori 130 minimum sample size to obtain a fit index, namely RMSEA, within a given range [0.00; 0.08] 131 (as recommended on the literature, Brown, 2015), with 90% power and $\alpha = .05$. Based on 132 simulations, the minimum sample size was N = 26 (data and code for this estimation are 133 available at https://osf.io/hpsjw/?view_only=6e28c8307294494e9eec45d2670efd8d) 134 (MacCallum et al., 2006b). Second, we estimated an a priori maximum sample size by using 135 the stopping rule based on resource constraints. Because we endorsed no priors regarding an 136 expected effect size, we decided to recruit as many participants as possible during the running 137 of the online study. During these three weeks, 535 people logged on. Exclusion criteria were 138 individuals who: (1) did not consent to participate; (2) completed the study multiple times; (3) 139 responded incorrectly to seriousness checks (Aust et al., 2013); (4) did not consider regular 140 physical activity and healthy diet as important for them (i.e., below 2 on a 1-7 goal-141 importance scale, Fishbach et al., 2003). In addition, we used the *performance* package 142 (version 0.10.0) (Lüdecke et al., 2021) for R-Studio (R Core Team, 2021) to identify 143 observations that were influential on the nine variables of interest (see Measures section). 144 Specifically, based on a composite score obtained via the application of multiple outlier 145 detection algorithms (Lüdecke et al., 2021), we excluded participants classified as influential 146 by at least half of the methods used by this package (the data and code of this data cleaning

147	are available at https://osf.io/hpsjw/?view_only=6e28c8307294494e9eec45d2670efd8d). The
148	application of these exclusion criteria in the described order led to a final sample size of $N =$

149 385 participants (154 women; $M_{age} = 19.42$, $SD_{age} = 2.71$).

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Finally, we conducted a sensitivity power analysis with G*Power 3.1.9.4 (Faul et al., 150 151 2007) to estimate the minimal effect size on the most constrained multiple regression of the 152 structural equation models. With our convenience sample of N = 385, n_{predictors} = 5 (see Measures), power = .90, α = .05, the smallest detectable effect size was f² = .04. This is one of 153 154 the smallest small-to-medium effect sizes (range = .02 to .15), in line with findings on the 155 relationships between self-control and health behaviors (Pfeffer & Strobach, 2018).

156 Measures

Independent Variables 157

Inhibition 158

159 A recent consensual stop-signal task² (Verbruggen et al., 2019) was used to assess 160 inhibition. The task included 216 trials in total, splitted in 3 blocks of 72 trials. A typical trial 161 started with a central fixation circle presented for 250 ms, followed by display of the stimulus (a right- or left-pointing white arrow within a circle) until the participant's response. The 162 instruction was to respond systematically according to the direction indicated by the arrow by 163 164 pressing a predefined keyboard button. However, participants had to stop their response (i.e., 165 not press the key) if a signal beep was made after the presentation of the arrow. The delay between the arrow's presentation and the beep was adjusted up or down by 50 ms as a 166 167 function of the participant's performance, starting with an initial delay of 250 ms. The delay

²We preferred the stop-signal task for inhibition because it has been proposed as more closely related to selfcontrol, as it captures more specifically behavioral control, compare to other task such as the Stroop task (supposedly more associated with cognitive inhibition) (Diamond, 2013). We preferred the letter-memory task for updating, because it was the most strongly associated with "updating" latent factor on Miyake et al., (2000) and Miyake & Friedman (2012). We preferred the number-letter task for updating, because it has been proposed as one of the more reliable, and as more appropriate for adults than other task (Diamond, 2013).

168 could be increased up to 1150 ms if the previous signal-stop was successful, and decreased 169 down to 50 ms if the previous signal stop failed. This delay is referred to as the Stop-Signal 170 Reaction Time (SSRT) and gives an estimation for response-inhibition latency in 171 milliseconds. We calculated the SSRT by using the integration method (Verbruggen et al., 172 2019). The lower the SSRT, the more difficult it is to stop the go-process, and the higher the 173 SSRT, the easier it is to stop the go-process. Accordingly, a lower (higher) SSRT integration 174 means the participant has stronger (weaker) inhibition.

175 Updating

We used a letter memory task to measure updating² (Friedman et al., 2008). A series 176 177 of letters appeared consecutively in the center of the screen for a duration of 2.5 s for each 178 letter. Written instructions asked participants to recall, in forward order, the last three letters 179 after the last letter's disappearance, by selecting the correct letters from a letter matrix 180 provided. They had to click "blank" if they skipped a particular letter. The number of letters per series varied randomly through time (5, 7, or 9 letters). In total, 12 measurement trials 181 182 were completed (four of each length). Answers were scored as correct even if the three letters 183 were not recalled in the correct order (Miyake et al., 2000). The more participants were able 184 to recall letters per trial, the better their updating was considered.

Switching 185

186 Switching was assessed with a number-letter task² (Miyake et al., 2000). This task 187 involves two categorization tasks, in which character pairs including a letter and a number 188 (e.g., 3T, 4A) were presented. The participants were asked to categorize the pair depending on 189 whether the letter was a consonant or vowel (i.e., letter task), or depending on whether the 190 number was odd or even (i.e., number task). The tasks alternated between categorizing rules 191 in a clockwise fashion, and thus used predictable location cues in a 2x2 matrix (i.e., the top of 192 the matrix for letter categorization and bottom of the matrix for number categorization). Odd-

numbered trials were set as "switch task" trials and even-numbered trials as "non-switch task"
trials. Participants responded by button press, and the next stimulus was presented 150 ms

195 after the response. The whole task was composed of 128 trials. The reaction time switch cost

196 was calculated by assessing the difference between the correct latency of switch trials and

197 non-switch trials (Miyake et al., 2000). A positive reaction time switch cost indicates a slower

198 response in switch trials, than in non-switch trials, and conversely. For example, a highly

199 positive reaction time switch cost indicates low switching. Trials with reaction times under

200 150 ms and above 2000 ms were excluded from analyses (Rossell & Nobre, 2004;

201 Schoonbaert et al., 2011).

202 Mediating Variables

193

194

203 Trait Self-Control

Trait self-control was assessed with the 13-item version of the Brief Self-Control Scale (Tangney et al., 2004). Participants responded to the following instruction: "For each sentence, choose what suits you best", on a 7-point Likert scale from 1 (*Don't agree at all*) to 7 (*Completely agree*), with regard to the different items ($\alpha = .80, \omega = .83$).

208 Self-Contr

Self-Control Resources

Self-control resources were assessed by the subjective vitality scale, as in previous studies (Forestier et al., 2018). Participants were asked to answer the 5-item questionnaire (e.g., "At the moment, I feel alive and full of vitality"), with the following instruction: "For each item, please indicate the general feeling you have experienced over the past 7 days, by selecting the most appropriate number" on a 7-point Likert scale from 1 (*Don't agree at all*) to 7 (*Completely agree*) ($\alpha = .89$, $\omega = .92$).

215 Dependent Variables

216

Physical Activity and Sedentary Activity

Physical activity and sedentary activity were measured using the International Physical
Activity Questionnaire (Craig et al., 2003). Participants were asked to indicate how much
time they had spent doing moderate-intensity and vigorous-intensity physical activities (i.e.,
MVPA), how much time they had spent walking, and how much time they had spent sitting
and/or lying down (i.e., sedentary activity) in minutes, in their daily life over the last 7 days.

222

Healthy and Unhealthy Diet

Healthy and unhealthy diets were assessed using the Healthy Eating Behavior Scale (Pelletier et al., 2004), composed of two subscales: four items related to a healthy diet (e.g., "I eat fruit and vegetables") and the remaining items related to an unhealthy diet (e.g., "I use white sugar"). Participants indicated their consumption frequency on a 7-item scale ranging from 1 (*once or twice per month*) to 7 (*more than three times per day*).

228

3. Results

229 Descriptive statistics and Pearson's correlation coefficients are provided in the 230 Supplementary Materials (Tables S1 and S2). Correlations' table revealed that self-control 231 resources are significantly correlated with trait self-control (r = .21), MVPA (r = .27), 232 sedentary time (r = -.14), healthy diet (r = .15), unhealthy diet (r = -.12) and switching (r = -.14)233 .13). It also stresses that trait self-control is only significantly correlated with healthy and unhealthy diet (r = .21 and r = -.27, respectively). The only behaviors significantly correlated 234 235 are MVPA and healthy diet (r = .25). Among executive functions scores, switching is 236 significantly corelated with MVPA (r = -.10), and inhibition with healthy diet (r = -.10).

Structural Equation Modelling 237

238 Analytical Strategy

239 The hypotheses were tested using structural equation modeling with the Lavaan package 240 (version 0.6-8, Rosseel, 2012) in R-Studio (RStudio Team, 2015) (the R script, raw data, and 241 be found Open Science Framework, analysis dataset can in the at 242 https://osf.io/hpsjw/?view_only=6e28c8307294494e9eec45d2670efd8d). We used a two-step approach (Anderson & Gerbing, 1988). The first step is a confirmatory factor analysis (CFA). 243 244 Hence, we verified the construct validity of the measurement model to estimate a reliable one, 245 by examining factor loadings, modification indices and model fit indices. Second, after a 246 satisfactory fit was achieved for the measurement model, we tested the structural model (i.e., 247 the hypothesized relationships between the variables). The results section presents only the 248 structural models (see Measurement Models in Supplementary Materials). Model fit was 249 assessed by examining the comparative fit index (CFI), Tucker-Lewis Index (TLI), and root-250 mean-square error of approximation (RMSEA), with a satisfactory model having a CFI and a 251 TLI over 0.90, a RMSEA below 0.05 (Brown, 2015). Finally, after an estimation of the full 252 hypothetical model, non-significant paths were removed to estimate model's parsimony reliability (MacCallum & Austin, 2000). The statistical significance was set at $\alpha = .05$. Data 253 254 were standardized prior to model estimation. Indirect effects (i.e., mediation) were estimated if 255 independent variables and mediators were significantly associated with the dependent 256 variables, with the *RMediation* package (version 1.2.0) (Tofighi & MacKinnon, 2011).

257

Physical Activity and Sedentary Activity Structural Model

258 The physical activity and sedentary activity structural model contained two latent 259 variables (trait self-control and self-control resources) and five observed variables (inhibition 260 score, updating score, switching score, MVPA score, and sedentary activity). Moreover, compared with the measurement model, we added covariances between the three executive 261

262	function scores. The full structural equation model yielded good model-fit indices (χ^2 (130) =
263	153.43, RMSEA = 0.02, 90% CI [0.00, 0.03], CFI = 0.99, TLI = 0.99). Partially confirming
264	our hypothesis (H1), only one of the three executive function scores considered was
265	significantly associated with self-control. Precisely, switching was negatively related to self-
266	control resources (β =10, 95% CI [18,01], p = .02, R^2 = .02), evidencing that the higher
267	the cost, the lower the self-control resources. Contrary to our hypotheses (H2, and H3),
268	executive function scores were not significantly associated with physical activity and
269	sedentary activity (i.e., no direct effect). In partial accordance with H4 and H5, only self-
270	control resources positively predicted MVPA ($\beta = .31, 95\%$ CI [.18, .43], p<.001, $R^2 = .08$)
271	and negatively predicted sedentary activity ($\beta =17, 95\%$ CI [30,05], $p = .01, R^2 = .03$).
272	Trait self-control was not significantly associated with healthy or unhealthy behaviors.
273	Mediation analysis revealed a significant indirect effect of switching on physical activity
274	(<i>indirect</i> =03, 95% CI [06,00], p = .02) and sedentary activity (<i>indirect</i> = .02, 95% CI
275	[.00, .04], p = .03), suggesting total mediations through self-control resources, which partially
276	confirm our hypothesis (H6). Figure 2 shows this structural model ³ .

277 Healthy and Unhealthy Diet Structural Model

The healthy and unhealthy diet structural model contained four latent variables (trait self-control and self-control resources, and healthy and unhealthy diet scores), three observed variables (i.e., inhibition, updating, and switching scores), and the covariances between the three executive function scores. The full structural equation model yielded good model-fit indices (χ^2 (188) = 218.35, RMSEA = 0.02, 90% CI [0.00, 0.03], CFI = 0.99, TLI = 0.99). Switching was still associated with self-control resources, as in the physical and sedentary activity model. Contrary to H2, and H3, there was no significant direct association between

³For exploratory purpose, we conducted the model separately on vigorous and moderate physical activity, and on walking. Tables of these models are in Supplementary Materials (S5, S6, S7)

285 executive functions with healthy and unhealthy diets. Mostly according with H4 and H5, 286 results revealed that trait self-control positively predicted healthy diet ($\beta = .28, 95\%$ CI [.07, .50], p = .01, R^2 = .10) and negatively predicted unhealthy diet (β = -.64, 95% CI [-.91, -.36], 287 $p < .001, R^2 = .19$), and that self-control resources also positively predicted healthy diet ($\beta =$ 288 .13, 95% CI [.02, .23], p = .02, $R^2 = .10$). Mediation analysis revealed a significant indirect 289 290 effect of switching on healthy diet (*indirect* = -.01, 95% CI [-.03, -.00], p = .02), suggesting a 291 total mediation through self-control resources. This partially confirms H6. Figure 3 shows this 292 structural model. Table 3 summarizes regression coefficients of the two structural models (the 293 complete tables are available in the Supplementary Materials, Tables S3 and S4).

294

4. Discussion

295 The present study tested the role of executive functions as predictors of individual 296 differences in trait self-control and self-control resources, which are likely also associated 297 with health behaviors (physical activity, healthy diet, sedentary activity, and unhealthy diet). Partially contrary to our hypotheses, we found small evidence supporting the role of executive 298 299 functions as predictors of individual differences in trait self-control or self-control resources. 300 Precisely, switching was only associated with self-control resources. Similarly, no direct role 301 of these functions regarding healthy and unhealthy behaviors was observed. Nevertheless, in 302 support of our hypotheses, there was an association between individual differences in self-303 control and health behaviors. Individuals with higher self-control resources practiced more 304 physical activity, ate healthier, and spent less time being sedentary than individuals with 305 lower self-control resources. It also indicated that individuals with higher trait self-control 306 adopted a healthier diet, and a less unhealthy one, than individuals with lower trait self-307 control. These findings stress the central role of self-control, and especially of self-control 308 resources, for adopting health behaviors. Finally, we observed three indirect effects of 309 switching on behaviors, totally mediated by self-control resources. Individuals with higher

310 switching were more physically active, less sedentary, and ate healthier, through higher self-311 control resources. These findings pave the way to a more refined understanding of the 312 predictors of individual differences in self-control resources, and to the ways by which 313 executive functioning could influence health behaviors.

314 Relations Between Executive Functions, Self-Control, and Health Behaviors

Partially according with Necka et al. (2018) and Saunders et al. (2018), our study did not find
strong evidence of significant associations between executive functions and self-control.
However, our intention was to examine the contribution of each executive function to selfcontrol and health behaviors. Indeed, it was proposed that updating and inhibition might be
the most important executive functions for self-control, enabling a better representation of the
goal and a better inhibition for fighting a threatening temptation (Hofmann et al., 2012).
However, our study does not support these suggestions.

Suggestions were divided concerning the role of switching. On the one hand, , when 322 323 facing a motivational conflict, individuals with high switching could efficiently switch from 324 an ineffective conflict resolution strategy (e.g., resisting temptation) to a situationally more 325 effective (e.g., avoiding temptation) (Hofmann et al., 2012). On the other hand, individuals 326 with high switching could present lower self-control as this switching could promote quick 327 and efficient disengagement from a goal-oriented mindset to a mindset oriented toward the 328 pursuit of tempting alternatives (Hofmann et al., 2012). Our results support partially the first 329 statement. Individuals with high switching are more likely to be able to shift their mind 330 toward subjective feelings permitting goal-attainment strategies, such as high self-control 331 resources, useful for efficient self-control act and adaptative motivational-conflict resolution.

Our findings also support the absence of evidence in favor of direct relations between
executive functions, trait self-control and physical activity, as observed in some previous
studies. Pfeffer and Strobach (2017) showed that most composite executive functions scores

335 they calculated were not significantly correlated with trait self-control, except for two 336 switching scores that were, modestly (i.e., task-cueing $R^2 = .04$, alternating-runs $R^2 = .03$; 337 Pfeffer & Strobach, 2017). They also revealed that most executive function scores were not 338 significant direct predictors of intention-behavior gap, except for updating score. These 339 findings highlighted no direct relations either between executive functions and trait self-340 control or between executive functions and a physical activity, which is consistent with the 341 current results. Interestingly, Pfeffer et Strobach (2017) found that half of the executive-342 function scores they considered (one inhibition, one updating, and one switching score) 343 moderated the effects of trait self-control on the physical activity intention-behavior gap. 344 Together with our results, this suggests that updating and inhibition are not direct predictors 345 of individual differences in self-control and physical activity, but could rather moderate the 346 relation between self-control and this behavior. However, because the other half of the 347 executive-function scores they considered (one inhibition score, one updating score, and one 348 switching score) showed no interaction with self-control to predict physical activity, further investigations are required. 349

350 Similarly, the fact that we found no significant direct relations between executive 351 functions and unhealthy diet were consistent with Hofmann et al. (2009). Indeed, there study 352 stressed no direct correlation between candy consumption and different components of 353 inhibition, but these latest (i.e., executive attention, behavioral inhibition, affect regulation) 354 consistently moderated the relations between automatic affective reactions and candy 355 consumption. All other things being equal, our study and the aforementioned results 356 (Hofmann et al., 2009; Necka et al., 2018; Pfeffer & Strobach, 2017) taken together suggest 357 that inhibition and updating are not direct predictors of self-control or healthy or unhealthy 358 behaviors, but could moderate relations between affective reactions (e.g., automatic affective 359 reactions, conscious experience of temptations) and behaviors, and between self-control and

health behaviors. Investigations of inhibition and updating as moderators of the relationbetween self-control and health behaviors would be of particular interest.

362 Relations Between Self-Control and Health Behaviors

363 All the relations found between self-control and health behaviors are consistent with the literature (de Ridder et al., 2011, 2012; Forestier et al., 2018). Specifically, self-control 364 365 resources were related to physical activity, sedentary activity, and healthy diet, while trait self-control was significantly related only to healthy and unhealthy diet. Furthermore, self-366 367 control aspects have different effects depending on the behavior examined, with quite a 368 similar effect size previously estimated (Forestier et al., 2018). Precisely, self-control 369 resources were related to health behaviors with a descriptively comparable effect size (i.e., R^2 = .08 for physical activity, and R^2 = .10 for healthy diet) and a smaller effect size for the 370 unhealthy behavior ($R^2 = .03$). It suggests that self-control resources are likely efficient to 371 372 promote health behaviors, but less effective in preventing the unhealthy ones. Health 373 behaviors show differences as well as similarities. For example, physical activity, and a 374 healthy diet are comparable as they require to be initiated, while sedentary activity and an 375 unhealthy diet are both things that need to be stopped (McEachan et al., 2010). Nevertheless, 376 physical activity requires more effort to be initiated than a healthy diet (McEachan et al., 377 2010), and sedentary activities appear to be attractive because they preserve energy 378 expenditure (Cheval & Boisgontier, 2021), while unhealthy food seems attractive because of 379 the immediate pleasure it provides (Appelhans, 2009; Volkow et al., 2011). Another 380 interesting perspective could be to investigate the influence of behavioral features on self-381 control aspects that could be effective or ineffective in the conflict resolution. This 382 encouraging finding requires further investigations, for example with other studies' design 383 (see "Strengths, Limitations, and Future Work Perspectives"), or with other instruments to 384 capture the behaviors, such as accelerometers for physical and sedentary activities, and daily

diet diary or questionnaires such as the new Healthy and Unhealthy Eating Behavior Scale
(Guertin et al., 2020) for diet behavior, to confirm the crucial role of self-control resources for
health behaviors.

388 Strengths, Limitations, and Future Work Perspectives

389 Several limitations of the present study need to be addressed. First, our study sample 390 consisted of young and relatively healthy students. Psychological determinants driving 391 behavior maintenance could differ from those driving behavior change (e.g., habit vs. coping 392 planning). Hence the current results need to be replicated in individuals engaged in health 393 behavior change processes. Second, the cross-section of the current study is insufficient to 394 understand the role of within-person variations in executive functions and self-control in daily fluctuations of health behaviors. Nevertheless, the within-person variations of inhibition seem 395 396 more predictive of snack consumption (i.e., unhealthy diet) than individual differences in 397 inhibition (Powell et al., 2017). In addition, an important part of the variance of self-control 398 resources is found at the within-person level (i.e., 60%) despite variance at the between-399 person level (Smolders et al., 2013). Thus, longitudinal study designs with daily repeated 400 measures will be required to properly examine the relationships of within-person executive 401 functions and self-control resource variations with health behavior fluctuations. Third, our 402 study considered executive functions and self-control as predictors of the overall level of 403 health behaviors over a week, without measuring the participant's intention to engage in these 404 behaviors. Despite including only individuals who considered physical activity and healthy 405 diet important for them, it remains possible that they did not support a particular intention to 406 engage in healthy behaviors during the week we considered. For example, some participants 407 may not have practiced physical activity because they did not intend to, rather than because of 408 low self-control. Future studies could assess intention-behavior gap instead of health 409 behaviors' global level over a week. In line with previous studies (Pfeffer et al., 2020; Pfeffer

410 & Strobach, 2017), participants could be asked their intention to engage in a certain quantity 411 of physical activity, to avoid a certain quantity of sedentary activity, to adopt a healthy diet, 412 and to avoid unhealthy food before and after all the measurements. Then the discrepancy 413 between intention endorsed and behaviors actually adopted (i.e., intention-behavior gap) 414 could be assessed to examine the role of executive functions and self-control in reducing this 415 gap. Finally, the self-reported nature of the self-control and behaviors measures is probably 416 another limitation of our study. As significant associations were observed between certain 417 self-reported measures (e.g., self-control resources and physical activity), we cannot exclude 418 that a part of these associations was because of a common method bias. However, we also 419 observed association between reaction-time task and self-reported measures (e.g., updating 420 and self-control resources), and also evidenced absence of significant association between 421 other self-reported measures (e.g., trait self-control and behaviors). Though, these specific 422 associations could rather suggest meaningful associations rather a method artifact. Altogether, 423 direct and conceptual (e.g., different measures and sample) replications could provide 424 valuable information, and identify the robustness of current (un)significant relations.

425 This study, nevertheless, has several strengths. We first examined the relative and 426 distinct role of the three executive functions, namely inhibition, updating, and switching, by 427 using executive tasks according to recent literature (e.g., Verbruggen et al., 2019). We also 428 tested our hypotheses with structural equation modeling, which increases the reliability of 429 scores and relations by (a) explicitly assessing the measurement error; (b) estimating latent 430 variable scores by scoring observed variables rather than other aggregating methods; and (c) 431 testing a model where a structure (e.g., covariances) could be imposed and assessed as to fit 432 of the data (Novikova et al., 2013). The final strength of our study is its good power to detect 433 small effect size (i.e., power = .90 for $f^2 = .04$).

Conclusion 434

No evidence was found supporting executive functions as direct predictors of the four 435 436 health behaviors considered. However, the current results support the role of self-control 437 resources as a potential way to promote physical activity and reduce sedentary activity, and 438 trait self-control as a likely determinant to increase adoption of a healthy diet and lower that 439 of an unhealthy diet. In sum, we hypothesize that *cold* executive functions may not explain 440 individual differences in self-control or health behaviors, and that the aspects of self-control 441 (trait or state) that are effective in health behavior adoption depend on the behavioral domain. 442 This study paves the way to longitudinal studies at the within-person level assessing the 443 effects of hot "affective-related" executive functions on trait and self-control resources and health behaviors. Those effects can be tested by using affectively-charged executive function 444 445 in which the neutral stimuli (e.g., arrows, letter or number) are replaced by affective stimuli such as pictograms of physical or sedentary activity, or of healthy of unhealthy food (e.g., see 446 447 Forestier et al., 2022b, p. 27-28, for a discussion). They can also be tested by using self-448 control measures specific to the behaviors considered (e.g., self-control resources for 449 resolving physical-activity related motivational conflict). 5. References 450 451 Allan, J. L., McMinn, D., & Daly, M. (2016). A bidirectional relationship between executive 452 function and health behavior: Evidence, implications, and future directions. Frontiers 453 in Neuroscience, 10(AUG), 1–13. https://doi.org/10.3389/fnins.2016.00386

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Table 1

	Physical Activity and	l Sedentary	y Time Model				
Independent variable	Dependent variable	Estimate	Estimate 95%CI [LL, UL]	Std. Err.	Z	р	R ²
Regression Slopes							
SSRT integration (I)	<u> </u>	0.02	[-0.06, 0.11]	0.04	0.58	.562	
Mean correct letters (U)	Self-control resources	0.02	[-0.06, 0.10]	0.04	0.47	.638	.015
RT switch cost (S)	resources	-0.10*	[-0.18, -0.01]	0.04	-2.27	.023	
SSRT integration (I)		-0.02	[-0.06, 0.03]	0.03	-0.61	.541	
Mean correct letters (U)	Trait self-control	0.00	[-0.04, 0.05]	0.02	0.18	.860	.001
RT switch cost (S)		0.00	[-0.04, 0.05]	0.02	0.20	.842	
SSRT integration (I)		0.04	[-0.06, 0.14]	0.05	0.81	.418	
Mean correct letters (U)		0.00	[-0.09, 0.10]	0.05	0.10	.924	
RT switch cost (S)	st (S) MVPA	-0.07	[-0.17, 0.03]	0.05	-1.44	.150	.081
Self-control resources		0.31***	[0.18, 0.43]	0.06	4.83	.000	
Trait self-control		0.06	[-0.18, 0.30]	0.12	0.51	.609	
SSRT integration (I)		-0.06	[-0.16, 0.04]	0.05	-1.11	.266	
Mean correct letters (U)		-0.01	[-0.11, 0.09]	0.05	-0.27	.790	
RT switch cost (S)	Sedentary activity	0.04	[-0.06, 0.14]	0.05	0.77	.441	.029
Self-control resources		-0.17**	[-0.30, -0.05]	0.06	-2.71	.007	
Trait self-control		-0.02	[-0.27, 0.23]	0.13	-0.16	.874	
Fit Indices							
χ^2		153.43					
χ^2_df		130.00					
p_{χ^2}		.079					
CFI TLI		.99 .99					
RMSEA		.02	[0.00, 0.03]				

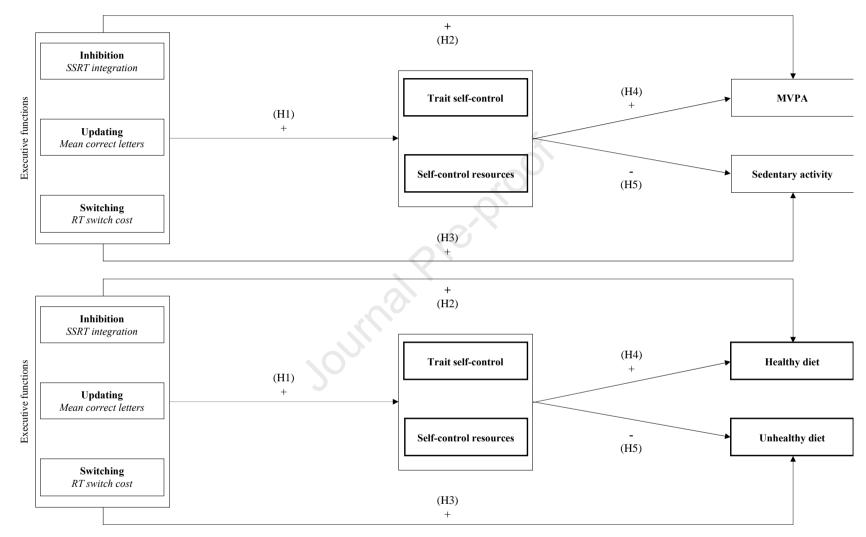
Regression Coefficients from Structural Equation Models

	Healthy and Unhealthy Diet Model					
Regression Slopes						
SSRT integration (I)	0.10	0.03	[-0.06, 0.11]	0.04	0.58 .559	
Mean correct letters (U) Self-control resources		0.02	[-0.06, 0.11]	0.04	0.49 .627 .016	
RT switch cost (S)		-0.10*	[-0.18, -0.01]	0.04	-2.28 .023	
SSRT integration (I)		-0.02	[-0.07, 0.03]	0.03	-0.61 .542	
Mean correct letters (U)	Trait self-control	0.00	[-0.05, 0.05]	0.03	0.18 .856 .002	
RT switch cost (S)		0.01	[-0.04, 0.06]	0.03	0.24 .813	
SSRT integration (I)		-0.07	[-0.16, 0.01]	0.04	-1.76 .078	
Mean correct letters (U)		0.07	[-0.01, 0.16]	0.04	1.73 .084	
RT switch cost (S)	Healthy diet	0.06	[-0.02, 0.14]	0.04	1.41 .159 .097	
Self-control resources		0.13*	[0.02, 0.23]	0.05	2.34 .019	
Trait self-control		0.28**	[0.07, 0.50]	0.11	2.62 .009	
SSRT integration (I)		0.02	[-0.07, 0.12]	0.05	0.47 .637	
Mean correct letters (U)		-0.02	[-0.11, 0.08]	0.05	-0.33 .743	
RT switch cost (S)	Unhealthy diet	-0.07	[-0.16, 0.03]	0.05	-1.36 .174 .194	
Self-control resources		-0.09	[-0.21, 0.02]	0.06	-1.55 .121	
Trait self-control	~0.	-0.64***	[-0.91, -0.36]	0.14	-4.56 .000	
	Fit I	ndices				
χ^2		218.35				
χ^2_df		188.00				
p_{χ^2}		0.06				
CFI		0.99				
TLI RMSEA		0.98 0.02	[0.00, 0.03]			

Note. *p<0.05, **p<0.01, ***p<0.001. MVPA = Moderate to vigorous physical activity, SSRT = Stop-Signal Reaction Time, I = Inhibition, U = Updating, S = Switching

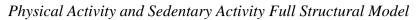
Figure 1

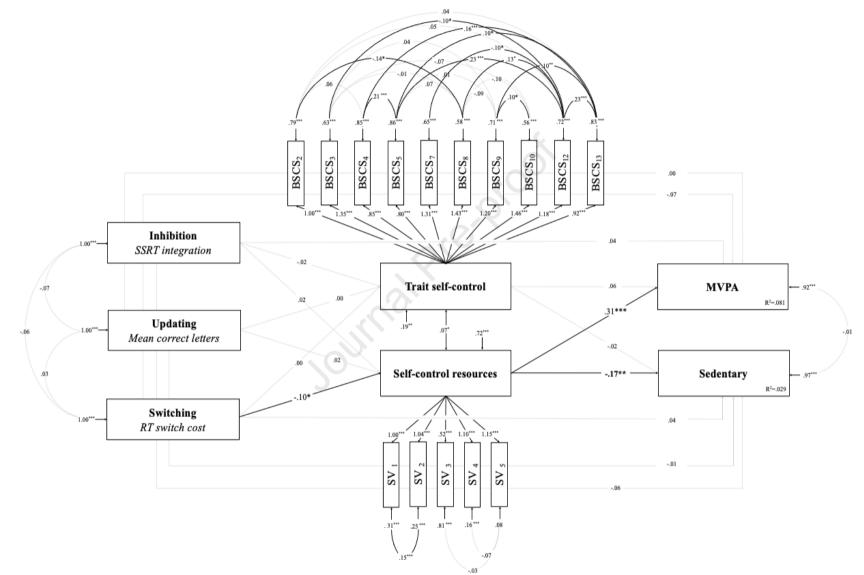
Hypothetical Models



Note. These are summarized hypothetical models. The full model will test the relations between each executive function, trait and state self-control and each behavior. RT = Reaction Time, SSRT = Stop-Signal Reaction Time, MVPA = Moderate to Vigorous Physical Activity. Variables with a bold line are latent. Variables with normal line are observed. H6 is the hypothesis related to the effects of executive functions on health behaviors mediated by self-control and is not illustrated

Figure 2

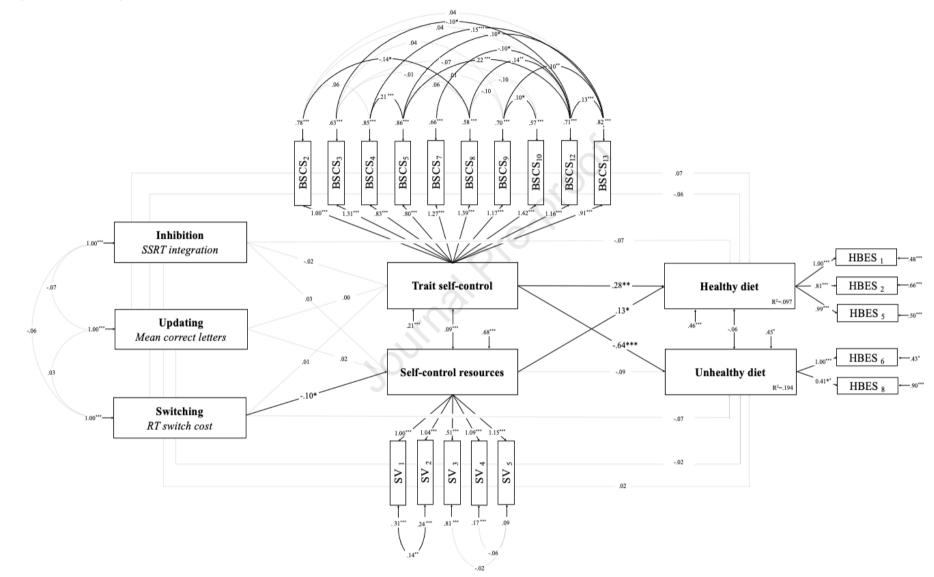




Note. *p<0.05, **p<0.01, ***p<0.001. SSRT = Stop-Signal Reaction Time, RT = Reaction Time, BSCS = Brief Self-Control Scale, SV = Subjective Vitality, MVPA = Moderate to Vigorous Physical Activity. Path darkness level distinguishes significant and non-significant relations.

Figure 3

Healthy and Unhealthy Diet Full Structural Model



Note. *p<0.05, **p<0.01, ***p<0.001. SSRT = Stop-Signal Reaction Time, RT = Reaction Time, BSCS = Brief Self-Control Scale, SV = Subjective Vitality, HBES = Healthy Eating Behavior Scale. Path darkness level distinguishes significant and non-significant relations.

- Trait self-control and self-control resources are associated with health behaviors
 - Executive functions could have an effect on self-control and health behaviors
 - Self-control resources predicted physical / sedentary activities and healthy diet
 - Trait self-control predicted healthy and unhealthy diet

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• Hot executive functions (affectively charged) seem promising for self-control

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The authors declare no conflict of interests.

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